Eccentric training for prevention of hamstring injuries may depend on intervention compliance: a systematic review and meta-analysis

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ABSTRACT
Background Hamstring injury is a prevalent muscle injury in sports. Inconclusive evidence exists for eccentric hamstring strengthening to prevent hamstring injuries. One reason for this discrepancy may be the influence of intervention non-compliance which has individual study estimates, and therefore pooled estimates.

Objective This systematic review aims to determine the effect of eccentric hamstring strengthening on the risk of hamstring injury and quantitatively explores the impact of intervention non-compliance on the precision, heterogeneity and strength of pooled estimates.

Methods A computer-assisted literature search of Medline, CINAHL, Cochrane, EMBASE, AMED, SportDiscus and PEDro databases was conducted with keywords related to eccentric strengthening and hamstring injury. The search was conducted from the end of a previous comprehensive review forward (1 December 2008 to 31 December 2013). Random-effects models were used for both main effects and a sensitivity analysis. Pooled estimate precision was measured with a confidence limit ratio (confidence limit ratio (CLR); upper limit divided by the lower limit) and heterogeneity was assessed with I2, Cochrane’s Q and τ2. A protocol was not registered for this review.

Results Four out of 349 studies met the inclusion criteria. In main effects analysis, eccentric hamstring training did not reduce the risk of hamstring injury (risk ratio [RR]=0.59 ([95% CI 0.24 to 1.44])). This estimate was imprecise (CLR=6.0) with significant heterogeneity (p value=0.02, 69.6% variation and τ2=0.57). Subjects compliant with eccentric strengthening had a significant (RR=0.35 ([95% CI 0.23 to 0.55])) reduction in hamstring injuries. This estimate was precise (CLR=2.4) and homogenous (p value=0.38, 2.8% variation and τ2=0.007).

Conclusions The null-biased effect in using intent-to-treat methods from intervention non-compliance has a substantial impact on the precision, heterogeneity and the direction and strength of pooled estimates. Eccentric strengthening, with good compliance, appears to be successful in prevention of hamstring injury.

INTRODUCTION
Hamstring strains are the most prevalent muscle injuries reported in sport. These injuries alone have accounted for between 6% and 29% of all reported injuries in football, Australian Rules football, rugby union, basketball, cricket and track sprinters. 1–9 Hamstring injuries also have a high recurrence rate. 2 10–12 Significant research effort over the past 10 years has been directed at injury prevention and return to sport after hamstring injury. 13–23 Continued high prevalence of injury rates from 2001 to 2008 24 suggest that traditional prevention and rehabilitation programmes have been ineffective. 25

Eccentric training for injury prevention has been suggested as a means of decreasing the high prevalence of these injuries, 13 19–22 with even more promising results than those observed with concentric training. 26 In fact, as pointed out by Thorborg, 27 the number needed to treat (NNT) of 13 to prevent 1 acute hamstring injury (new or recurrent) is considerably more impressive than the NNT of approximately 90 for neuromuscular training programmes for the prevention of anterior cruciate ligament injury. 28 While the precise mechanism by which eccentric hamstring exercise is effective is not fully understood, 29 and continues to be debated, 29 30 the potential benefit of eccentric training has been recognised. 27

Clinical trials commonly implement intention to treat (ITT) analyses to compare end points regardless of postrandomisation compliance. 31 This type of analysis tends to underestimate the true effect when participants are not compliant and the treatment is effective. 32 Failure to comply can be especially problematic with ITT methods because as compliance decreases, effectiveness also decreases regardless of the efficacy of the treatment. 32 Non-compliance rates as high as 20 and 22% using ITT methods have produced substantial bias in estimating treatment effects. 33 Commonly, investigators will recognise non-compliance and report per-protocol, or analysis of the ‘complier’ groups separate. 34 This type of subgroup analysis can also be problematic since the baseline risk of compliers and non-compliers may be different, which has the potential to create a selection bias that overestimates the true effect. 35 Decreased compliance with eccentric training for the prevention of hamstring injuries has been recognised in some, but not all, reported randomised trials 20 21 and is suspected to be one reason for drastically differing results. 19–22 These differing results have led to inconclusive evidence to support eccentric hamstring strengthening to prevent hamstring injury in pooled analysis.

The purpose of this review is twofold: (1) systematically review the literature on randomised clinical trials (RCT) examining the effect of eccentric hamstring strengthening on prevention of hamstring injury among athletes, and (2) quantitatively explore the impact of intervention compliance on the precision, heterogeneity and both strength and direction of pooled study estimates.

METHODS

Study design
The design of this systematic review was developed using the guidelines of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA). The PRISMA statement includes a 27-item checklist that is designed to be used as a basis for reporting systematic review of randomised trials. A review protocol was not registered for this review.

Search strategy
Identification and selection of the literature
A systematic, computerised search of the literature in PubMed, CINAHL, EMBASE, Cochrane, AMED, SportDiscus and PEDro was conducted by a research and education librarian (LL) with controlled vocabulary and keywords related to eccentric strengthening and hamstring injury. This coauthor did not participate in aspects of screening, full-text review or data abstraction. Our search time-frame was restricted from 1 December 2008 to 31 December 2013 due to a previous comprehensive review also investigating this topic from 1966 to December 2008. The PubMed search strategy used for the current study can be found in our online supplementary appendix. We developed our search strategy from a previous comprehensive review’s search strategy using updated Medical Subject Headings (MeSH) terms, included terms specific to eccentric strengthening and included terms to search for any systematic reviews or meta-analysis published that may be similar to our work. The search strategy for this previous comprehensive review can be found in the online supplementary appendix of their publication. The reference lists of all selected publications were checked to retrieve relevant publications that were not identified in the computerised search. References of screened and included articles, abstracts and available conference proceedings were also hand searched by one of the authors (MPR) and included publications, posters, abstracts or conference proceedings. To identify relevant articles, two reviewers independently screened titles and abstracts of all identified citations (MPR and APG). Full-text articles were retrieved if the abstract provided insufficient information to establish eligibility or if the article had passed the first eligibility screening.

Selection criteria
All articles examining eccentric injury prevention for the hamstrings were eligible for full-text review. An article was eligible for study inclusion if it met all of the following criteria: (1) the article was a RCT; (2) included athletes (participation in organised sports) of either sex who were at risk of incurring hamstring injuries and not participating in a hamstring rehabilitation programme; (3) athletes with and without a history of hamstring injury; (4) included eccentric exercise intervention in comparison with a control or alternative interventions for the prevention of hamstring injuries; (5) study was required to report on the outcome of incidence of hamstring injury and 6) compliance with the intervention was either reported or could be calculated.

An article was excluded if: (1) included athletes with existing, or under treatment for, lower-limb musculoskeletal injuries; (2) reports focused on children below the age of 10 years, or (3) the article was not in English.

All criteria were independently applied by two reviewers (MPR and APG) to the full text of the articles that passed the eligibility screening of titles and abstracts. Any disagreement was resolved by discussion.

Quality assessment
The methodological quality of individual studies was assessed using the Physiotherapy Evidence Database (PEDro) scale. Results from individual study analysis of quality were used to identify common areas of methodological weaknesses across studies. Two reviewers (MPR and LH) independently assessed the quality of the studies. Disagreement or ambiguous issues were resolved by discussion with an independent assessor (APG).

PEDro uses 11 criteria, and reviewed studies were awarded one point for each criterion that was clearly satisfied, for a potential maximum value of 11 points. Criteria included: (1) eligibility criteria reported; (2) random assignment; (3) concealed allocation; (4) groups similar at baseline regarding most important prognostic indicator; 5) blinding of participants; (6) blinding of therapists who administered the therapy; (7) blinding of assessors who measured key outcome; (8) measures of at least one key outcome were obtained from more than 85% of initial participants; (9) all participants received treatment or control condition as allocated; (10) results of between-group statistical comparisons are reported and (11) study provides point measures and measures of variability for at least one key outcome.

Statistical analysis
Data, including counts and risk or rate ratios with corresponding CIs, were extracted manually from included studies by two separate authors (LD and AK) and confirmed by a third author (MPR). DerSimonian and Laird random effects models were used for all pooled analyses, which incorporate both between-study and within-study heterogeneity, to produce summary pooled risk ratios (RRs) and 95% CIs. We used this model to be consistent with a previously reported review on the same intervention and outcome. Heterogeneity test statistics and their p values were used to assess consistency of reported RRs across studies. I-squared statistic (I^2) was used to describe the percentage of total variation across studies due to heterogeneity rather than chance alone with values of >50% to indicate substantial heterogeneity. Significant heterogeneity was indicated with a p<0.10. A higher p value was chosen to test for heterogeneity since these tests have low power particularly when there are few studies analysed. The τ^2 is reported to describe the pooled among-study variance of true effects, thereby reflecting the magnitude of heterogeneity. Precision of individual and overall pooled study estimates was determined with a confidence limit ratio (CLR). This ratio is the upper confidence limit divided by the lower confidence limit with results closer to “1” indicating greater relative precision. Publication bias was not formally assessed due to low statistical power with few included studies. All analyses were conducted in Stata 12 (Stata Corp, College Station, Texas, USA) by one of the authors (APG).

Sensitivity analyses
A simple sensitivity analysis was conducted to compare the treatment effects analysing estimates from compliant subjects (ie, per protocol analyses). These analyses represent the effect of eccentric hamstring strengthening on hamstring injury with good intervention compliance. Estimates and 95% CIs were abstracted from articles reporting analysis of complier groups separately. These estimates and 95% CI’s were entered into our statistical software. This software recalculates the SDs resulting in slight differences in the 95% CIs. The magnitude of difference is less than a 100th of a decimal point. DerSimonian and Laird random effects models were employed in the same
fashion as the main effects analysis above. Measures and criteria for heterogeneity and precision were also identical to above. Both complier analysis and ITT individual study estimates were qualitatively compared and the overall pooled estimates between the main effects and the sensitivity analysis were compared with a change-in-estimate formula ($|\ln RR_{\text{main effects}} - \ln RR_{\text{sensitivity analysis}}|$).

RESULTS

Selection of studies

Three hundred forty-nine titles were identified through database and reference searches. Eight abstracts then were assessed for eligibility for inclusion, resulting in four appropriate studies (1229 athletes).\textsuperscript{19–22} Three of the four full text articles screened for inclusion were excluded due to inability to determine either study compliance and/or intervention outcomes.\textsuperscript{44–46} Only one additional study\textsuperscript{22} was discovered since the previous published review.\textsuperscript{36} The number of athletes in the studies ranged from 30\textsuperscript{19} to 942\textsuperscript{22}. Three of the studies\textsuperscript{19–21} investigated various levels of soccer players, and the fourth study\textsuperscript{20} investigated Australian Rules football players. One study\textsuperscript{19} utilised a “Yo-Yo” flywheel ergometer, while the other three studies\textsuperscript{20–22} implemented the Nordic hamstring exercise as the method of eccentric hamstring strengthening. The process of study selection and the number of studies excluded at each stage, with reasons for exclusion is available in figure 1. A description of the included studies populations, interventions, comparators and outcome measures are found in table 1.

Quality of studies

The scores on each of the 11 criteria and total scores for each study are presented in table 2. One study\textsuperscript{22} scored 9/11, second study\textsuperscript{19} scored 8/11, third study\textsuperscript{20} scored 5/11 and the fourth study\textsuperscript{21} scored 3/11 on PEDro scoring.

Summary of individual studies

Table 3 provides the reported compliance, adverse events, individual study estimates and measure of precision for each of the included articles. Reported compliance was poor in the Gabbe\textsuperscript{20} and Engebretsen\textsuperscript{21} studies, with 50% of enrolled athletes declining to participate\textsuperscript{20} and 63% reporting not being compliant with any exercise\textsuperscript{21} respectively. Both the Askling et al\textsuperscript{19} and Petersen et al\textsuperscript{22} studies reported good compliance with over 90% of the athletes having completed the intervention. Delayed onset muscle soreness (DOMS) was by far the primary reason for adverse event reporting. The individual study estimates from the Askling et al\textsuperscript{19} and Petersen et al\textsuperscript{22} studies were identical, indicating a 70% reduction in hamstring injuries from eccentric

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Figure 1  Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) flow diagram for study inclusion.
<table>
<thead>
<tr>
<th>Study</th>
<th>Athletes</th>
<th>Interventions</th>
<th>Comparator(s)</th>
<th>Number of injuries</th>
<th>Outcome recording</th>
</tr>
</thead>
<tbody>
<tr>
<td>Askling et al</td>
<td>30 male soccer athletes from 2 premier league teams in Sweden. Goalkeepers, injured players and players with chronic hamstring problems were excluded from the study</td>
<td>“Yo-Yo” flywheel ergometer (n=15). 16 sessions of specific hamstring length training, every 5th day for the first 4 weeks, and every 4th day during the last 6 weeks (total of 10 training weeks)</td>
<td>The training and control groups (n=15) followed the same protocol with the only exception being that the training group received additional specific hamstring training during phase 1. Intervention was for 10 weeks period</td>
<td>Intervention group: 3/15 athletes&lt;br&gt;Control group: 10/15 athletes</td>
<td>Compliance and side effects: (muscle soreness/symptoms in training) recorded individually by athlete. Injuries: recorded by team training staff (not part of study)</td>
</tr>
<tr>
<td>Gabbe et al</td>
<td>220 Australian Rules Football athletes. Athletes were excluded if they had a hamstring injury at the time of recruitment</td>
<td>Nordic Hamstring Exercise (n=114). 12 sets×6 reps with a 10 s rest in between reps and 2–3 rest between sets. 5-session programme over a 12-week period. The first three sessions were 2 weeks apart during the last 6 weeks of preseason. The final two sessions were 3 weeks apart during the first weeks of the season</td>
<td>Intervention group: 10/114 athletes&lt;br&gt;Control group: 8/106 athletes</td>
<td>Intervention group: 10/114 athletes&lt;br&gt;Control group: 8/106 athletes</td>
<td>Compliance and side effects: Each session was supervised by study personnel. Injuries: monitored by club officials and medical staff. Injury surveillance form filled out/injury recorded by team physiotherapist not blinded to allocation of player to specific group</td>
</tr>
<tr>
<td>Engebretsen et al</td>
<td>37 male soccer athletes from the 1st, 2nd, or top of the 3rd Norwegian League. Athletes divided into high (previous injury) and low (no previous injury) risk groups</td>
<td>Nordic Hamstring exercise (n=24). Week 1: (1 session per week with 2 sets of 5 reps), week 2: (2 sessions per week with 2 sets of 6 reps), week 3: (3 sessions per week with 3 sets of 6–8 reps), week 4: (3 sessions per week with 3 sets of 8–10 reps), weeks 5–10: (3 sessions per week with sets of 12, 10, and 8 reps)</td>
<td>Control group (n=13) continued with their normal training schedule throughout the 10-week cycle</td>
<td>Intervention group: 5/24 athletes&lt;br&gt;Control group: 1/13 athletes</td>
<td>Compliance and side effects: Individual players recorded exercises performed. Injuries: team physical therapist instructed exercises and recorded injuries (blinding unknown, but assumed they were not since they provided intervention, etc.)</td>
</tr>
<tr>
<td>Petersen et al</td>
<td>942 Danish professional and amateur soccer athletes. Detail of previous hamstring injuries during past 12 months was recorded</td>
<td>Nordic Hamstring Exercise (n=461). Along with normal training performed 27 sessions of Nordic hamstring exercise over 10 week span. Week 1: (1 session per week with 2 sets of 5 reps), week 2: (2 sessions per week with 2 sets of 6 reps), week 3: (3 sessions per week with 3 sets of 6–8 reps), week 4: (3 sessions per week of 8–10 reps), week 5–10: (3 sessions per week with 3 sets of 12–10–8 reps). After week 10 the players continued to do 1 session per week of 3 sets of 12–10–8 reps</td>
<td>Control group (n=481). Continued with their normal training schedule throughout the 10-week cycle</td>
<td>Intervention group: 15 (12 new and 3 recurrent)/461 athletes&lt;br&gt;Control group: 52 (32 new and 20 recurrent)/481 athletes</td>
<td>Compliance and side effects: Supervised by team coaches. Injuries: team physiotherapist recorded injuries. Recurrence of previously recorded injuries in registration period was not included to avoid recording same injury more than once</td>
</tr>
</tbody>
</table>
training with the Peterson et al \(^22\) study having greater precision due to the increased sample size. Gabbe et al \(^20\) and Engebretsen et al \(^19\) both reported increased risk of hamstring injuries from eccentric training that were not statistically significant and substantially imprecise.

### Meta-analysis findings

Individual study estimates reporting ITT analysis, overall pooled estimate and measures of heterogeneity are illustrated in figure 2. This figure represents each included study’s reported RR and 95% CI of hamstring injury following eccentric training. The diamond represents the overall summary estimate and the width of the diamond represents the overall point estimate 95% CI. The overall pooled estimate from the main effects analysis was 0.59 (95% CI 0.24 to 1.44) with a CLR of 6. Significant heterogeneity (heterogeneity p value=0.02) was found with 69.6% of the variation in the overall pooled RR attributable to heterogeneity and \(\tau^2=0.542\).

Figure 3 illustrates the findings from the sensitivity analysis that included individual study estimates with ITT estimates from Asklings et al \(^19\) and Peterson et al \(^22\) and per protocol analysis from both Gabbe et al \(^20\) and Engebretsen et al \(^21\). These pooled studies were also homogeneous (heterogeneity p value=0.38) with 2.8% of the variation in the overall pooled RR attributable to heterogeneity and \(\tau^2=0.007\). The overall pooled estimate was 0.35 (95% CI 0.23 to 0.53) with a CLR of 2.39 indicating a 65% reduction in the risk of hamstring injury from eccentric hamstring strengthening.

There was a substantial increase in the strength of the overall pooled estimate from the sensitivity analysis when compared with the main effect analysis (ie, 0.59 to 0.35) that was down and away from the null value. This difference represented a 52% change-in-estimate from the main effect to the sensitivity analysis.

### DISCUSSION

This systematic review assessed the literature on eccentric strengthening as a prevention measure of future hamstring injury and explored the effect that compliance may have on the precision, heterogeneity and both the strength and direction of pooled study estimates. Our search yielded four studies \(^19\)-\(^22\) that met our inclusion/exclusion criteria. Using data from these sources we found a stronger effect towards injury prevention than the previously reported review. \(^16\) Similar to this previous

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**Table 2** The PEDro quality assessment of individual studies

<table>
<thead>
<tr>
<th>Journal article</th>
<th>PEDro scale</th>
<th>Score</th>
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<tbody>
<tr>
<td>Askling et al (^19)</td>
<td>+ + + + + + + + + +</td>
<td>8</td>
</tr>
<tr>
<td>Engebretsen et al (^21)</td>
<td>+ + + + + + + + + +</td>
<td>5</td>
</tr>
<tr>
<td>Gabbe et al (^20)</td>
<td>+ + + + + + + + + +</td>
<td>6</td>
</tr>
<tr>
<td>Petersen et al (^22)</td>
<td>+ + + + + + + + + +</td>
<td>9</td>
</tr>
</tbody>
</table>

1. Eligibility criteria were specified.
2. Subjects were randomly allocated to groups.
3. Allocation was concealed.
4. Groups were similar at baseline regarding most important prognostic indicators.
5. Blinding of therapists who administered the therapy.
6. Blinding of all assessors who measured at least one key outcome.
7. Measures of at least one key outcome were obtained from more than 85% of the participants.
8. All subjects for whom outcome measures were available received the treatment or control condition as allocated.
9. Results of between-group statistical comparisons are reported for at least one key outcome.
10. Study provides both point measures and measures of variability for at least one key outcome.
+ met criteria; − criteria not met.

PEDro, Physiotherapy Evidence Database.

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**Table 3** Reported compliance, adverse events, individual study estimates and precision of included studies

<table>
<thead>
<tr>
<th>Study</th>
<th>Reported compliance</th>
<th>Reported adverse events (number of events)</th>
<th>ITT individual study estimates RR (95% CI)</th>
<th>Precision (CLR)</th>
<th>Per protocol individual study estimates RR (95% CI)</th>
<th>Precision (CLR)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Asklings et al (^19)</td>
<td>100%</td>
<td>DOMS (n=11 athletes)</td>
<td>0.30 (0.10 to 0.88)</td>
<td>8.8</td>
<td>NA</td>
<td>NA</td>
</tr>
<tr>
<td>Gabbe et al (^20)</td>
<td>Intervention group: 50% decline in participation rate from session 1 to session 2 Control group: 10% decline in participation rate from session 1 to session 2</td>
<td>DOMS was primary reason reported for lack of compliance (n=NR)</td>
<td>1.15 (0.48 to 2.83)</td>
<td>5.9</td>
<td>0.30 (0.08 to 1.12)</td>
<td>14.0</td>
</tr>
<tr>
<td>Engebretsen et al (^21)</td>
<td>Only 21.1% (12 athletes) completed 20 or more training sessions 62.2% of players reported not having performed any exercises, 7.9% of athletes performed 1–9 sessions and 7.9% of athletes performed 10–19 sessions</td>
<td>DOMS (n=NR)</td>
<td>2.71 (0.35 to 20.79)</td>
<td>59.4</td>
<td>0.94 (0.29 to 3.07)</td>
<td>10.6</td>
</tr>
<tr>
<td>Petersen et al (^22)</td>
<td>Team compliance was an average of 91% of the intended training sessions performed during the 10-week midseason period (8% dropout rate in intervention group and 9% dropout rate in control group) Compliance was not measured after the first 10 weeks</td>
<td>DOMS (n=&quot;most players in the intervention group&quot;)</td>
<td>0.30 (0.17 to 0.53)</td>
<td>3.1</td>
<td>NA</td>
<td>NA</td>
</tr>
</tbody>
</table>

CLR, confidence limit ratio; DOMS, delayed onset muscle soreness; NA, not applicable; NR, not reported; RR, risk or rate ratio.
review, our pooled estimate in main effects analysis was not statistically significant. Our sensitivity analysis found that those studies with a large proportion of non-compliance explained the decreased precision and heterogeneity in the main effects analysis.

One additional study by Petersen et al.\textsuperscript{22} has been published on this topic since the 2010 Cochrane Review.\textsuperscript{16} This study differed from the other three included studies in that it utilised a cluster (team) randomised approach; however, the intraclass correlation was taken into account during the analysis. Despite the majority of information (weight) provided by this study to our overall pooled estimate, the addition of the Petersen et al.\textsuperscript{22} study alone continues to support that there is inconclusive evidence for eccentric hamstring strengthening in the prevention of hamstring injury. Our heterogeneity analysis, with the inclusion of this additional study,\textsuperscript{22} continued to indicate significant heterogeneity between studies, which is also consistent with this previous Cochrane Review.\textsuperscript{16} The addition of the Petersen et al.\textsuperscript{22} study, however, did allow for interesting comparisons in our sensitivity analysis to explore the impact of compliance on the overall pooled estimate and heterogeneity.

Our sensitivity analysis yielded several important findings. First, the inclusion of estimates of compliant groups from the Engebretsen et al.\textsuperscript{17} and Gabbe et al.\textsuperscript{20} studies, together with the estimates of compliant groups from the Askling et al.\textsuperscript{19} and Peterson et al.\textsuperscript{22} studies, yielded a pooled estimate that indicates substantial (65%) decreased risk of hamstring injury from eccentric training. Interestingly, this is a similar finding to that of Arnason et al.\textsuperscript{13} who also reported a 65% reduction in hamstring injury from eccentrics in an observational intervention study. The estimate of subjects compliant with the intervention from Gabbe et al.\textsuperscript{20} was identical to the findings of Askling et al.\textsuperscript{19} and Peterson et al.\textsuperscript{22} Similarly, the estimate of the compliant group from Engebretsen et al.\textsuperscript{17} shifted towards a decrease in hamstring injury when compared with the ITT estimate from the same study. These are important considerations since one of the main concerns of using complier group or per protocol estimates are that it may lead to an overestimation of the

Figure 2  Forest plot reflecting individual study intent-to-treat analysis estimates and 95% CI, an overall pooled estimate (diamond) and 95% CI with heterogeneity statistics (I\textsuperscript{2} and heterogeneity p value).

Figure 3  Forest plot of individual study estimates reflecting good intervention compliance and 95% CI, an overall pooled estimate (diamond) and 95% CI and heterogeneity statistics (I\textsuperscript{2} and heterogeneity p value).
Improving exercise compliance is multidimensional. Addressing the issues outlined above is likely to improve compliance prior to and during an injury prevention study. Implementation of coaching workshops, policy enactment by sport organisations due to demonstrated benefit of such programmes, emphasising performance-enhancement benefits to the athlete and focusing on athletes at risk of injury for each specific sport are additional means of improving injury prevention exercise compliance.

In addition to non-compliance, a concern of the studies in this review was the method of randomisation and the potential risk of contamination. Three of the studies randomised the athletes within the same team, leading to the potential for contamination. A potentially bigger risk of contamination was the fact that 19 of the 31 teams in the Engberg et al study did team-based preventive exercises similar to those employed in the study on a regular basis throughout the pre-season. The potential for athletes in the control group to have also shown benefit from prevention training must be considered in this study. Only the Petersen et al study randomised by team, potentially leading to social support within the team as a positive compliance factor. Although the potential risk of contamination exists it is unlikely to result in the substantial between-study estimate differences especially in the presence of significant non-compliance.

Our review is not without limitations. Although we conducted a thorough search utilising multiple databases, our search was limited to the English language, potentially excluding relevant studies. There are concerns with including estimates from compliant subjects (i.e., per protocol) along with ITT estimates in the sensitivity analysis as these estimates of compliant subjects may overstate the true effect. However, in this particular case, there appears to be considerable consistency when analysing estimates reported from studies with good compliance. Moreover, our review found a substantial improvement in the statistical homogeneity and relative precision of the overall pooled estimate in the sensitivity analysis with compliant groups only. Other factors related to the reporting of non-compliance may also lead to bias and influence individual study and pooled estimates; however, we were unable to assess these details finely in these analyses. Analyses of groups with good compliance, in this particular case, appear to reflect the effect of eccentric hamstring exercise on the prevention of hamstring injury among those athletes performing the exercise rather than effectiveness of the delivery method represented in ITT methods.

CONCLUSION
Intent-to-treat is the preferred method of determining effectiveness of interventions in RCT’s. However, this method is subject to null-bias when there is substantial non-compliance. When non-compliance is recognised within studies, authors would benefit from assessing the impact of compliance by analysing compliant groups separately, especially in the presence of differential non-compliance between study arms. These estimates may be compared with prior studies to determine if non-compliance is producing a substantial null-bias effect. In producing overall pooled estimates with studies examining the effect of eccentric exercise on hamstring injury, intervention non-compliance appears to produce a substantial null-bias effect. Non-compliance, therefore, may have important implications in meta-analysis conclusions and recommendations for implementation of eccentric strengthening for hamstring injury prevention. Eccentric strengthening with good compliance appears to be an influential factor in successful prevention of hamstring injury.
injury. Future studies and training programmes should focus on the implementation of eccentric strength training with an emphasis of successful compliance.

**Contributors**

APG, MPR, LH, LD, AK, DB, LL, CP and AT planned, conducted and reported the described intervention compliance. Both APG and MPR contributed equally to the writing of this manuscript. Both APG and MPR contributed equally to the writing of this manuscript.

**Competing interests**

None.

**Provenance and peer review**

Not commissioned; externally peer reviewed.

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Appendix. Literature search used in study